

1 the Salinas Valley between late-January and early-June in 2013 and 2014. Levels of *P. effusa*
2 DNA were determined by a species-specific qPCR assay. *Peronospora effusa* was detected prior
3 to and during the growing season in both years. Nonlinear time series analyses on the data
4 suggested that the within-season dynamics of *P. effusa* airborne inoculum are characterized by a
5 mixture of chaotic, deterministic, and stochastic features, with successive data points somewhat
6 predictable from the previous values in the series. Analyses of concentrations of airborne *P.*
7 *effusa* suggest both an exponential increase in concentration over the course of the season and
8 oscillations around the increasing average value that had season-specific periodicity around 30,
9 45, and 75 days, values that are close to whole multiples of the combined pathogen latent and
10 infectious periods. Each unit increase in temperature was correlated with 1.7-6% increased odds
11 of an increase in DNA copy numbers, while each unit decrease in wind speed was correlated
12 with 4-12.7% increased odds of an increase in DNA copy numbers. Disease incidence was
13 correlated with airborne *P. effusa* levels, and a receiver operating characteristic curve analysis
14 suggested that *P. effusa* DNA copy numbers determined from the spore traps nine days prior to
15 disease rating could predict disease incidence.

17 The shift towards pre-washed, bagged spinach has increased the demand for fresh market
18 spinach production in the US. Over 85% of the total US spinach production occurs in California
19 and Arizona (USDA-NASS 2015). The most serious threat to spinach production is downy
20 mildew, caused by the obligate biotrophic pathogen, *Peronospora effusa* (Grev.) Rabenh
21 (formerly *P. farinosa* f. sp. *spinaciae*) (Choi et al. 2007). Severe downy mildew can result in
22 post-harvest rotting of spinach in salad bags or packages, leading to reduced quality and shelf life
23 of bagged products (McKay et al. 1992). Additionally, severe spinach downy mildew may cause
24 consumers to be wary of visibly blemished products because of previous food poisoning

1 outbreaks, and the potential for harboring foodborne pathogens in leaf lesions (Arnade et al.
2 2009; Simko et al. 2015). Consequently, packing companies routinely reject spinach lots at low
3 levels of downy mildew, with anecdotal reports of disease incidence thresholds as low as one
4 percent.

5 Approximately 40% of the annual US fresh market spinach crop is produced in the Salinas
6 Valley in California, a highly productive agricultural region responsible for supplying a large
7 proportion of many other leafy green vegetables in the US (USDA-NASS 2015). The valley is
8 approximately 150 km long and 40 km wide, and experiences strong diurnal fluctuations in
9 temperature, relative humidity, and wind caused by a subsidence inversion, more commonly
10 referred to as a marine layer. These conditions typically cause a favorable environment for fog
11 and low clouds to develop with accompanying leaf wetness. At times, the inversion effect can be
12 sufficiently strong that it resists thinning and breakup of low clouds. The resulting high relative
13 humidity and lack of surface heating help to maintain leaf wetness well beyond the typical burn
14 off time of mid-morning (approximately 10 AM). The marine layer is a persistent regime during
15 the spring to early fall months in the Salinas Valley, and causes the valley to remain much cooler
16 than surrounding foothills that are above the marine inversion, even during middays and
17 afternoons in the early summer.

18 Spinach growers rely heavily on host resistance and fungicides to manage downy mildew.
19 This reliance on resistance was relatively effective until recently, when virulent forms of *P.*
20 *effusa* began to emerge in rapid succession (Correll et al. 2011; Irish et al. 2007). Organic
21 spinach producers are especially vulnerable to these virulent strains because synthetic fungicide
22 use is prohibited and cultivar choice is determined at planting. This has led to significant yield
23 losses in organic spinach production.

1 Conventional and organic spinach growers could potentially benefit from advanced warning
2 about disease pressure to plan optimal spray regimen and planting dates, respectively. Many
3 effective epidemiological models have been developed to predict plant diseases based on
4 environmental conditions that are conducive to growth and reproduction of the pathogen
5 (Dewdney et al. 2007; Gent et al. 2013; Gilles et al. 2004; Peduto et al. 2013). One potential
6 source of error in weather-based models is the implicit assumption of the presence of pathogen
7 inoculum (Carisse et al. 2008). While inoculum is often present in many systems, uncertainty
8 about the presence/absence of pathogen inoculum in specific locations and times can lead to
9 over-application of pesticides in the absence of the pathogen. Monitoring for the presence of
10 airborne spores or other pathogen inoculum can help increase the accuracy of disease forecasting
11 models (Carisse et al. 2008). Many downy mildew pathogens spread rapidly through the
12 production of wind-blown asexual sporangia. While most of the asexual spores remain local to
13 the epidemic, some of these spores can travel long distances and can initiate epidemics in other
14 regions or countries (Fontaine et al. 2013; Lamondia and Aylor 2001; Ojiambo et al. 2011;
15 Schmale and Ross 2015; Wu et al. 2001).

16 Recently, several studies have combined the use of spore-trapping devices and disease
17 forecasting models for devising regional integrated pest management programs. The integration
18 of these tools with other disease risk indicators often increases the accuracy of the disease
19 forecasting models by assessing the local presence of the pathogen rather than assuming
20 presence. Detection methods have significantly advanced in recent years, from passive grease-
21 covered glass slides to solar-powered impaction spore trap samplers linked with quantitative
22 PCR (qPCR) methods, enabling detection of very low numbers of spores (Carisse et al. 2008;
23 Carisse et al. 2009; Fallacy et al. 2007; Klosterman et al. 2014). These advances have made it

1 possible to accurately predict disease pressure in some cases, leading to more effective and
2 efficient disease control measures (Carisse et al. 2011). Use of spore trapping technologies can
3 help properly time control strategies, and increase basic understanding of pathogen biology and
4 epidemiology (Urbez-Torres et al. 2010).

5 One reason for the recent shift towards species-specific qPCR detection is the co-occurrence
6 of several species with morphologically similar structures (Fallacy et al. 2007; Klosterman et al.
7 2014). Recent studies have shown that the organisms that cause downy mildew on plants in the
8 Amaranthaceae are phylogenetically distinct from one another, and often associated with specific
9 hosts (Choi et al. 2015). While there are some morphologically distinct traits that can distinguish
10 between these species (Choi et al. 2007), many related organisms have nondescript
11 morphologies, making accurate microscopic identification difficult. Klosterman et al. (2014)
12 developed *P. effusa*-specific primers that can distinguish between the downy mildew pathogen
13 on spinach and the closely related *P. schachtii* that infects beet and Swiss chard. During the
14 development of this assay, DNA of the spinach downy mildew pathogen was consistently
15 detected throughout the Salinas Valley, even in warm and dry southern portions of the valley.
16 This result was surprising given that downy mildew diseases typically develop in cool, moist
17 environments.

18 This study had the following objectives. First, to characterize the temporal pattern of
19 airborne inoculum of *P. effusa* in the Salinas Valley during the critical period of the growing
20 season from March through June. Second, to examine the degree of association between weather
21 variables and the observed pattern of airborne inoculum. Third, to examine the strength of
22 association between measured airborne inoculum concentration and observed disease incidence
23 in spinach crops.

MATERIALS AND METHODS

1
2 Spore trapping and *Peronospora effusa*-specific qPCR. The rotating arm impaction spore
3 trap samplers used in this study were obtained from Dr. Walt Mahaffee (United States
4 Department of Agriculture –Agricultural Research Service [USDA ARS], Corvallis, OR). The
5 traps were deployed as previously described (Klosterman et al. 2014) with the exception that
6 they were placed in pairs, at a spacing of one to two meters between traps at four sites in the
7 Salinas Valley (Fig. S1). The sites were selected for their proximity to spinach production areas.
8 The traps were deployed in late January in both 2013 and 2014, prior to the start of spring
9 spinach plantings in early-March. The 40 mm stainless steel, grease-coated spore trap rods
10 (316LSi stainless steel; Harris Products Group, Mason, OH) were collected three times a week
11 from late-January to early-June in both years and stored at 4°C until processing. The rods were
12 processed within two weeks after collection and remaining samples were stored at -80°C.

13 DNA extractions from the impaction spore trap sampler rods (NucleoSpin® Plant II kit;
14 Machery-Nagel) and qPCR (TaqMan) assays with the SNP-specific primers for each species, *P.*
15 *effusa* and *P. schachtii*, were carried out and results were expressed as copy number of *P. effusa*
16 ribosomal DNA as previously described (Klosterman et al. 2014).

17 Disease incidence rating. Spinach fields located near traps were rated for disease incidence
18 weekly using a cluster sampling method. Fields were divided into 70 m × 70 m sections and
19 rated using a notional ‘X’ pattern to separate individual sampling points, which were located at
20 the arm-ends of the ‘X’. Four adjacent 1 m² quadrats of spinach were rated at each sampling
21 point for disease incidence (presence or absence on individual leaves) of downy mildew
22 symptoms by estimating the proportion of infected leaves in each quadrat. Samples of diseased

1 leaves were collected from affected fields and DNA from *P. effusa* and the qPCR assay was used
2 to cross-validate the visual symptoms.

3 Weather data. High quality weather data for individual sites were generated by Fox Weather,
4 LLC using the MtnRT® custom software described in a summary by Fox (2011). Hourly
5 interpolation of meteorological variables including temperature, relative humidity, wind speed,
6 and wind direction on a 1.5 km grid was done by MtnRT for both seasons, with supplemental
7 data from the virtual weather stations as available (Coop et al. 2009).

8 Analysis: Log-normal distribution: The spore trap data were log-transformed (natural
9 logarithms) after first removing zero values, and the remaining data were compared with a
10 randomly formed dataset drawn from a log-normal distribution with the same mean and variance
11 as the data using the Kolmogorov–Smirnov (KS) test.

12 Non-linear time series analysis: The overall dynamical properties of the spore trap data
13 series from each site-year were assessed using NLTSM software (Nonlinear time-series
14 modelling, <http://www.eeb.uconn.edu/people/turchin/NLTSM.htm>; Turchin 2003). In particular,
15 the coefficient of prediction and Lyapunov exponent values were used to indicate the balance
16 between stochastic noise and chaos in the series (Turchin 2003).

17 Trend and periodicity: Linear regression of the natural log of the spore trap copy numbers
18 against time was used to extract the trend from each series. The resulting residuals were fitted to
19 a period-variable periodic wave function of form

$$20 \quad y = \sin((2\pi/\theta)t)$$

21 where θ is the period, y are the residuals of the linear regression, and t is the time. The value of θ
22 was increased in 1-unit increments from 1 to 1 less than the length of each series and the
23 coefficient of determination (R^2) was recorded. The periodic function with the highest R^2 value
24 was fitted to the data, and the fitted values subtracted from the data to recover a second set of

1 residuals, which were, in turn, assessed for evidence of periodicity. This iterative process of
 2 recovering statistically significant periods was repeated until the residuals contained no
 3 detectable periodicity. A statistical process model for each year's data was constructed by
 4 combining the estimated linear trend for the mean trap DNA copy number with the significant
 5 periods identified for that season. The fitted process model had the form

$$6 \quad y(t) = a + bt + \sum_n \sin\left(\frac{2\pi}{\theta_n}\right)(t)$$

7 where $y(t)$ is the estimated logarithm of the trap DNA copy number at time t , a is the intercept of
 8 the linear regression, b is the slope year-specific linear rate of increase with time, and θ_n are the n
 9 period values detected in each year's data.

10 Logistic regression of spore trap data on weather variables. Spore trap data were transformed
 11 by calculating differences between successive observations. This allowed transformation of the
 12 series of n observations into series of $n-1$ binary values, in which a value of 1 corresponded to a
 13 pair of successive observations between which *P. effusa* DNA copy number increased and 0
 14 corresponded to a period in which copy number decreased. Weather variables from the trapping
 15 periods were divided into six four-hour periods per day and logistic regression was used to
 16 examine the potential for weather data averaged within periods to predict the probability of
 17 observing increases in *P. effusa* DNA copy number.

18 Receiver operating characteristic curve. The spore trap data were assessed for their
 19 predictive capacity for future disease incidence of nearby crops. A case/control approach was
 20 used (Yuen and Hughes 2002, Carisse et al. 2008). Cases were defined as location-times at
 21 which average site disease incidence rating was greater than or equal to 0.1% and controls were
 22 defined as crops with less than 0.1% incidence by site. Spore trap data at successive lag times
 23 before each disease observation were translated into an ordinal categorical scale using e^x as the

1 scale values, where x is a member of the set of values 0, 0.5, 1, 1.5...13. Each trap observation
2 was allocated to the closest scale value. The predictive capacity of the traps was assessed by
3 taking each scale value in turn and assigning predictions of future disease status as either at-or-
4 above or below the disease incidence threshold of 0.1%, depending on the spore trap index value.
5 The predicted status of disease at each location-time as a case or control could then be compared
6 with its actual case/control status obtained from the disease sampling exercise; the sampling data
7 acted as the gold standard for comparison. Overall accuracy of the trap data at different
8 combinations of time lag and scale values was assessed by calculating Youden's Index (= true
9 positive proportion + true negative proportion - 1) and the positive and negative likelihood ratios
10 for cases for each lag/scale value combination (Hughes and McRoberts 2014). The positive
11 likelihood ratio for cases is the weight of evidence contained in a trap reading that a future
12 observation of disease will be above the threshold, while the negative likelihood ratio measures
13 the weight of evidence against a future observation of disease being above the threshold (Yuen
14 and Hughes 2002; Carisse et al. 2008; Hughes and McRoberts 2014). The combined data set for
15 2013 and 2014 consisted of 87 crop location-times, composed of 46 gold standard cases and 41
16 gold standard controls.

17 **RESULTS**

18 In all site-years, there were days in which there was no *P. effusa* detectable on either of the
19 traps in a pair (Fig. 1). When these events were removed, the distribution of *P. effusa* detected
20 approximately fit a lognormal distribution (Fig. 2) according to a Kolmogorov–Smirnov test,
21 although in 2013 there was a significant departure from the lognormal distribution as a result of
22 left skew in the data.

1 *Peronospora effusa* was detected throughout both growing seasons at all four sites (Fig. 1
2 and Fig. 4). On average, the detectable *P. effusa* levels were higher in 2013 than in 2014.
3 Results from the NLTSM analysis suggested that most of the spore trap time series had a
4 Lyapunov exponent that was negative or close to zero. Two of the nine site-year time series
5 (Soledad in 2014 and King City South in 2013) had positive Lyapunov exponents. The
6 coefficient of prediction for all site-years was consistently above zero (Fig. 3). These results are
7 indicative of a data-generating process that has some short-term predictability, but has
8 considerable stochasticity.

9 The mean DNA copy number per trapping interval was 1509 in 2013 and 796 in 2014.
10 Across all four sites, there was an overall increase in the spore copy number in both years as the
11 season progressed. When transformed to a natural log scale, the increasing trend over time for
12 spore copy number was described using simple linear regression in both years (Fig. 4).

13 After removal of the increasing time trend, results from fitting a periodic function to the
14 residuals of the linear regression suggested that there was a significant periodicity in both 2013
15 and 2014 (Fig. 4). In 2013, significant periodicity occurred at intervals of 26 and 44 days, while
16 in 2014, significant periods of 30, 45 and 75 days were detected (Table 1). Combining the linear
17 trend with periodic forcing (sin) functions for each year resulted in statistical process models that
18 accounted for approximately 83% of the variance in trap DNA copy number in 2013 and 61% in
19 2014 (Fig. 4).

20 The logistic regression analysis of the relationship between changes in airborne inoculum
21 levels and weather variables suggested that there was a higher probability of observing an
22 increase in trap DNA copy number when temperatures were higher than average during any of
23 the four-hour periods examined (Table 2). In addition, decreased wind for all time periods,

1 except between 6 AM – 10 AM, was significantly correlated with higher probability of observing
2 an increase in *P. effusa* DNA levels on traps. Increased relative humidity in the periods 10 AM –
3 2 PM and 6 PM – 10 PM had a low but significant positive correlation with increased probability
4 of observing higher *P. effusa* levels. Similar effects were noted for the weather conditions one
5 lagged trapping period before each observation of change in the DNA copy number.

6 The patterns of disease incidence and spore copy number levels followed each other in both
7 years (Fig. 5). Disease incidence levels were higher on average in 2013 compared with 2014,
8 matching the situation for spore copy numbers. In 2013, disease incidence increased
9 dramatically in April and peaked twice, in mid- to late-April and again in mid- to late-May. In
10 2014, the maximum levels of disease were lower than in 2013, but disease appeared earlier in the
11 season (in mid-March) and had three peaks, mid- to late-March, mid-April and early- to mid-
12 May. *Peronospora effusa* DNA extracted from pathogen samples collected from infected fields
13 was consistently amplified using the primers developed in Klosterman et al. (2014), suggesting
14 that the primers are robust.

15 Results from the receiver operating characteristic (ROC) curve analysis suggested that spore
16 copy number is able to predict disease incidence levels at 0, 1, and 4 lags prior to the disease
17 incidence rating, roughly approximating lags of 0, 2, and 9 days. Observations corresponding to
18 lag 4 gave the highest overall predictive accuracy and are the only ones among those with
19 predictive capacity that would allow sufficient time for a response by growers. At lag 4, using a
20 spore trap copy number of e^8 as the action threshold, Youden's Index was 0.35. The positive
21 likelihood ratio for disease increase was 2.36 and the negative likelihood ratio for disease
22 increase was 0.53. Observing e^8 trap copy numbers or more on day t more than doubled the odds
23 of observing greater than 0.1% disease at day $t+9$, while observing fewer than e^8 copies on traps

1 on day t reduced the odds of observing 0.1% or more disease on day $t+9$ by roughly half. There
2 was little to no predictive capacity of the spore trap data at 2 and 3 time lags for subsequent
3 disease incidence, as indicated by the proximity of the ROC curves for these lags to the line of
4 no discrimination (Fig. 6).

5 DISCUSSION

6 In this study, levels of *P. effusa* were monitored by impaction spore trap samplers and qPCR
7 at four sites throughout the Salinas Valley of California, a major spinach production region in the
8 US. The overall objectives of this study were to determine whether the spinach downy mildew
9 pathogen, *P. effusa*, is detectable prior to and during the spring growing season in the Salinas
10 Valley, whether levels vary between locations, and whether pathogen DNA copy number levels
11 are related to local disease incidence or weather variables. Understanding how airborne
12 inoculum concentration (as estimated by trap spore copy number) varies within and between
13 locations can help understand basic aspects of the pathogen biology as well as adjust practical
14 aspects of disease management (Carisse et al. 2009; Carisse et al. 2011). Understanding how
15 local disease levels and weather variables affect spore load is also a vital part of the assessment
16 of whether the technology can be used to improve disease management by growers.

17 When zero values were removed from the DNA copy number datasets, the resulting data
18 were approximately normal on a logarithmic scale. The lognormal distribution is commonly
19 observed in many biological systems, including other studies monitoring airborne microbial
20 populations (Limpert et al. 2001). The occurrence of this distributional form in the data can be
21 explained as the maximum entropy outcome for measurements of ratios (Frank and Smith 2011),
22 and reflects the way that qPCR estimates the sample DNA quantities by using cycle numbers
23 required to generate a threshold response level. The relative under-representation of data in the

1 lower tail in 2013 may reflect stochastic sampling effects, which result in the traps failing to
2 catch any inoculum by chance when the mean concentration is low, combined with the limit of
3 detection for the qPCR.

4 Long term dynamics of *P. effusa* levels can help describe underlying processes in the disease
5 system. The Lyapunov exponent (λ) provides a diagnostic for whether systems are stochastic
6 and will converge to an equilibrium ($\lambda < 0$) or chaotic and will diverge from an equilibrium ($\lambda > 0$)
7 (Turchin 2003). The results from the NLTSM analysis suggest that the spinach downy mildew
8 pathosystem in the Salinas Valley lies at the border between these dynamic domains, with the
9 balance between chaotic and stochastic properties in specific time series depending on location
10 and year. These observations are consistent with the idea that many pathogen population
11 dynamics are driven by both exogenous and endogenous factors that combine stochastic noise
12 with broad deterministic rules.

13 In general, endogenous life history factors (such as reproduction) for microbes are strongly
14 influenced by exogenous, stochastic variation in environmental conditions. Discrete events
15 introduced by human actions such as crop harvest or pesticide application, in contrast, can
16 impose local discontinuities in population dynamics, leading to the generation of chaotic
17 properties (Hughes and Gonzalez-Andujar 1997). Many real populations are thus subject to a
18 mixture of noise-generating and chaos-generating influences, and display dynamics in the
19 transition zone between the two domains; a phenomenon that Turchin (2003) characterized as
20 quasi-chaotic dynamics.

21 While the Lyapunov exponent values suggest that the pathosystem is close to the boundary
22 between stochastic and chaotic dynamics, the coefficient of prediction was consistently greater
23 than zero. This suggests that the dependence of *P. effusa* airborne inoculum levels on past levels

1 can be estimated and that such a statistical description will provide a better predictive estimate
2 than simply using mean levels (Turchin 2003). The presence of statistically significant trends
3 and periodic components in the data from both years supports the implications from the NLTSM
4 analyses. These analyses are relatively novel in plant pathology, although Skelsey et al. (2009)
5 used Turchin's (2003) coefficient of prediction as one of several metrics to evaluate
6 spatiotemporal models of potato late blight disease. The combination of the coefficient of
7 prediction and Lyapunov coefficient has been used to describe dynamics in other biological
8 systems, notably several insect pests (Turchin 2003). Plaza et al. (2012) found that the dynamics
9 of plant populations were also influenced by both exogenous and endogenous factors. They
10 noted that while some of their populations seemed to be driven by endogenous factors that might
11 be expected to generate chaos, very few field studies have been confirmed as chaotic systems.

12 *Peronospora effusa* detection increased exponentially over the course of growing seasons
13 both in 2013 and 2014, consistent with other spore trapping studies (Byrne et al. 2000, 2005;
14 Carisse and Pillion 2002; Granke et al. 2011). In the case of an obligate biotroph such as *P.*
15 *effusa*, it is probable that the exponential increase in airborne inoculum is driven by host
16 availability in the Salinas Valley. Host availability dramatically increases after the initial spring
17 planting in March each year and varies with market demand and individual crop growth duration
18 until June, when crop production temporarily shifts to alternative growing areas.

19 Although *P. effusa* was detected almost continuously throughout both growing seasons, there
20 was strong evidence for intra-season periodicity. The dominant periods differed between 2013
21 and 2014, but stable periodicity was observed in both years at approximately 30 and 45 days. In
22 2014, a longer period of 75 days occurred. It is currently not clear what the source of the
23 periodicity is. It may arise as an emergent property of the interaction between overlapping

1 cycles of infection and factors connected with the cropping cycles. The latent period for spinach
2 downy mildew is thought to be about 7 days, and the infectious period is about 9 days (van den
3 Bosch et al. 1988).

4 The observed dominant periodicities in the data correspond to approximately two, three, and
5 five times the combined latent-infectious period for the pathogen, suggesting that peaks in spore
6 trap detection may have followed periods when planting duration was sufficient to allow
7 multiple rounds of the pathogen infection cycle. The multiple asynchronous infection cycles and
8 asynchronous planting and harvesting schedules likely distort any natural periodicity in spore
9 production arising from the intrinsic life cycle biology of the pathogen. The observed dynamics
10 of the time series of trap catches are consistent with this hypothesis.

11 *Peronospora effusa* was detectable at all sites more or less continuously throughout both
12 growing seasons in 2013 and 2014. In comparable studies of other plant pathogens, detection of
13 airborne inoculum has been found to be more sporadic (Carisse et al. 2009; Eskalen et al. 2013;
14 Granke et al. 2009). Spinach is often grown year round, with a smaller acreage of winter crops
15 grown for freezer processing or for local markets, and with overall denser plantings in the
16 southern portion of the valley. Overhead irrigation is a routine practice, providing ample
17 moisture for the sporulation and splash dispersal of many fungal and oomycete pathogens
18 (Eskalen et al. 2013; Savory et al. 2011; Urbez-Torres et al. 2010). Moreover, the northern end
19 of the valley is regularly inundated with a marine fog layer. This regular presence of fog and the
20 dense constant cropping may allow the spinach downy mildew pathogen to maintain some level
21 of sporulation year-round, generating a background level of airborne inoculum. The presence of
22 a background inoculum level complicates the use of spore traps as a decision support tool, since

1 it removes the possibility of using the traps as pathogen detection devices and basing
2 management decisions on the basis of a simple presence/absence of pathogen inoculum.

3 During all observation periods, there was a significant association between warmer weather
4 and increased detection of pathogen DNA, similar to what was found by Granke et al. (2014) in
5 cucurbit downy mildew. These observations might be due to sporangiophores drying and
6 releasing their spores more in warmer conditions. In addition, there was a consistent, significant
7 effect from lower wind speeds and increased DNA detection, with the exception of the 6-10 AM
8 time period. This may be due to reduced sporulation in more windy conditions, which has been
9 observed under controlled conditions in the analyses of lettuce downy mildew (Scherin et al.
10 1995). While relative humidity plays a large role in many other downy mildew pathosystems
11 (Gilles et al. 2004), it did not seem to have a large effect on increased detectable levels of *P.*
12 *effusa* in our study. It is possible that microclimate effects on relative humidity near infected
13 fields may influence sporulation, and that these effects might not be detectable using somewhat
14 localized measurements of relative humidity from individual weather stations.

15 The ROC analysis suggested that the spore traps had predictive value for disease incidence as
16 far as 4 lags prior to disease rating, or roughly 9 days in advance. This number is just longer
17 than the reported latent period of 7 days for spinach downy mildew (van den Bosch et al. 1988),
18 suggesting that the spore traps may be detecting deposition events. The statistical dependence
19 between spore trap data at day t and disease incidence at day $t+9$ will be modulated by the
20 weather conditions during the interval between the prediction date, influencing the probability of
21 inoculum detected at day t infecting a host crop and the rate of pathogen growth in the interval (t ,
22 $t+9$). While controlled environment studies are necessary to confirm the exact environmental

1 requirements for infection and development of *P. effusa*, we analyzed our field data for weather
2 variables that correlate with increased *P. effusa* levels.

3 In theory, monitoring for the presence of a pathogen can significantly improve disease
4 management and can help in understanding aspects of its life cycle, which cannot be directly
5 observed. Our study discovered that *P. effusa* is consistently present throughout the Salinas
6 Valley, *P. effusa* populations increase both exponentially and with a periodicity over the course
7 of single growing seasons, and that weather variables can be correlated with airborne *P. effusa*
8 levels. In addition, airborne inoculum levels provide some advance warning of subsequent risk
9 of detectable disease in the crop. Ultimately, we hope that these tools can be used to help
10 growers predict and counteract disease with timely application of fungicides at reduced cost.

11 **ACKNOWLEDGMENTS**

12 We thank the California Leafy Greens Research Program (CLGRP) and the California
13 Department of Agriculture Specialty Crop Block Grant Program (Number SCB14043) for
14 funding this research. We thank Dr. Walt Mahaffee (USDA-ARS, Corvallis, OR) for providing
15 some spore trap materials and advice, and Lorena Ochoa (USDA-ARS, Salinas, CA) and Ruben
16 Pena (Hartnell College, Salinas, CA) for collecting spore trap samples.

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1 TABLE 1. Parameter estimates for fitting linear+periodic statistical models to spore trap DNA
 2 copy number data for *Peronospora effusa* in two growing seasons in the Salinas Valley in
 3 California. Data were converted to natural logarithms before parameter estimation

Year	Parameter	Parameter Estimate	Standard Error Estimate	t-Value (prob)
2013	Intercept	-0.072	0.4942	-14.54 (<2E-16)
2013	Time coefficient	0.046	0.0031	14.70 (<2E-16)
2013	Periodic (26 days)	-0.447	0.1640	-2.73 (<0.009)
2013	Periodic (44 days)	-0.604	0.1715	-3.52 (<0.001)
2014	Intercept	-0.043	0.7359	-5.84 (<4E-07)
2014	Time coefficient	0.020	0.0060	3.30 (<0.002)
2014	Periodic (30 days)	0.767	0.2300	3.33 (<0.002)
2014	Periodic (45 days)	0.745	0.2350	3.17 (<0.003)
2014	Periodic (75 days)	-1.290	0.2309	-5.59 (<9.5E-07)

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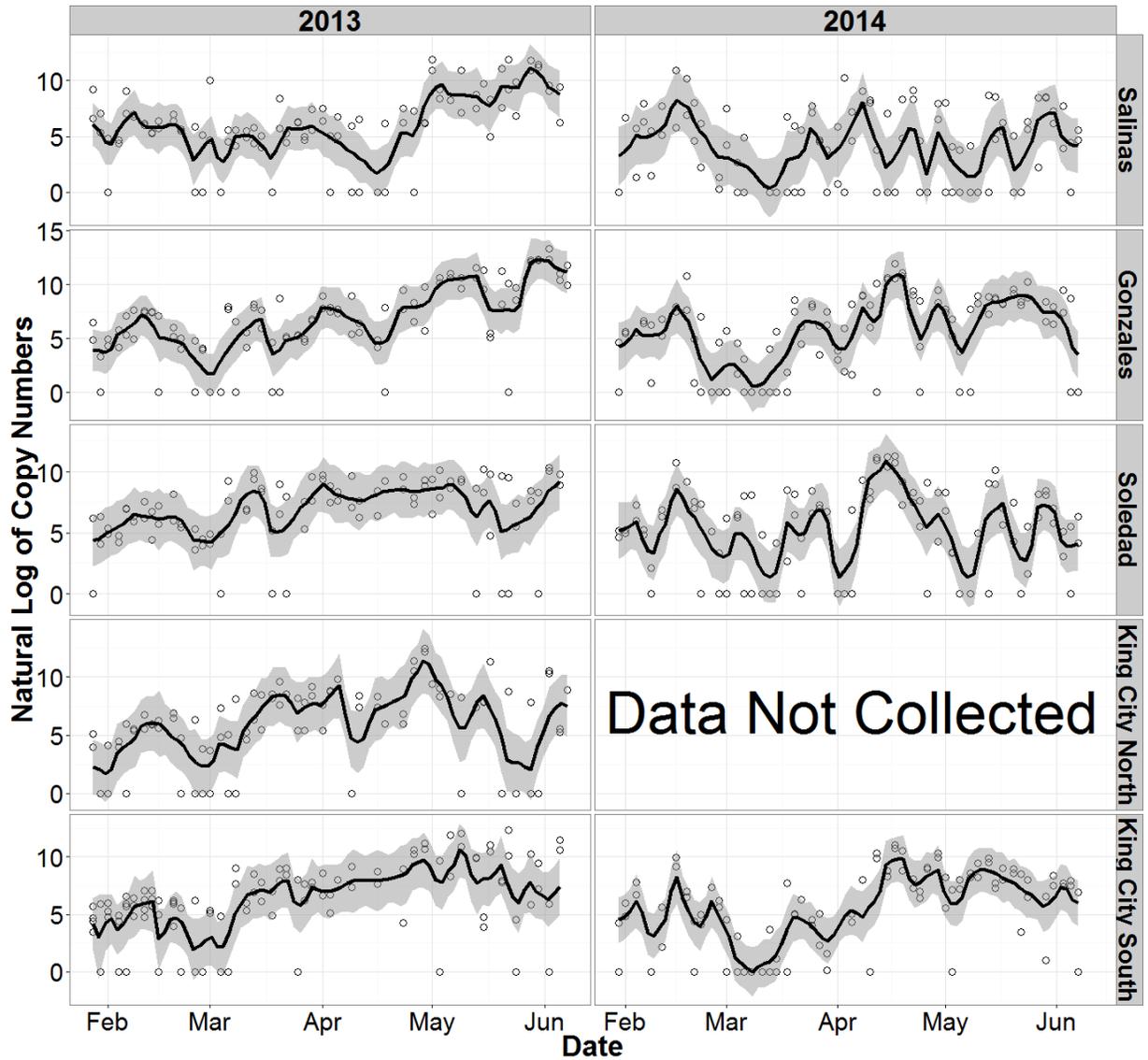
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1 TABLE 2. Exponentiated coefficient results of the logistic regression of ribosomal DNA copy number change by different
 2 environmental factors.

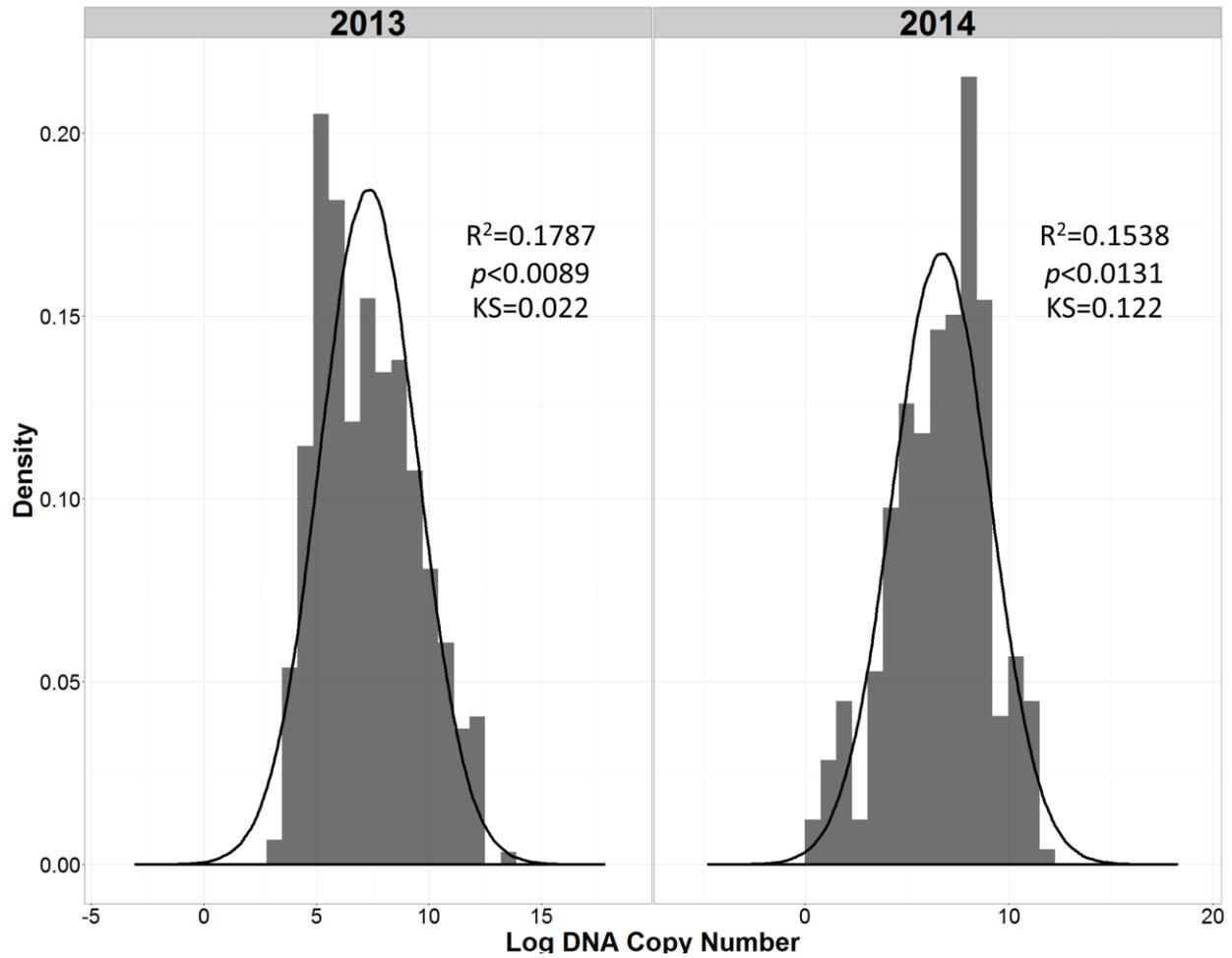
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Lag	Environmental Factor	10 AM - 2 PM		2 PM – 6 PM		6 PM – 10 PM		10 PM – 2 AM		2 AM – 6 AM		6 AM – 10 AM	
0	Air Temperature (°C)	1.0422	<0.0001	1.0421	<0.0001	1.0595	<0.0001	1.0392	0.0002	1.0339	0.0003	1.0167	0.0430
0	Relative Humidity (%)	1.009	<0.0001	1.0044	0.0625	1.0054	0.0250	1.0036	0.1778	1.0046	0.1634	1.0043	0.1083
0	Wind Speed (m/s)	0.9606	0.0405	0.9094	<0.0001	0.9404	0.0006	0.8733	<0.0001	0.8995	0.0012	0.9884	0.6779
1	Air Temperature (°C)	1.0203	0.0201	1.0233	0.0113	1.0359	0.0007	1.0235	0.0213	1.0009	0.9183	1.0099	0.2258
1	Relative Humidity (%)	1.0072	0.0016	1.0029	0.2128	1.0041	0.0846	1.005	0.0594	1.0095	0.0046	1.0072	0.0072
1	Wind Speed (m/s)	0.949	0.0073	0.9397	<0.0001	0.9601	0.0223	0.9425	0.0308	0.9565	0.1697	0.9606	0.1539



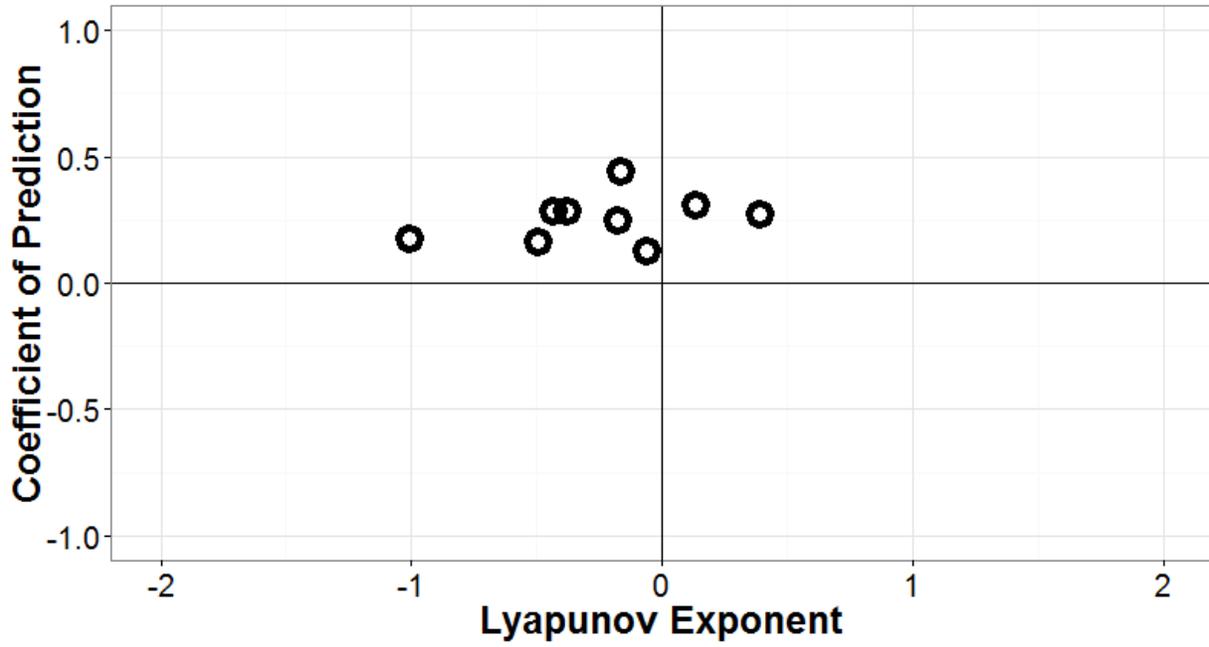
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 2 **Fig. 1:** Natural log transformed ribosomal DNA copy numbers of *Peronospora effusa* from the 9
 3 site-years. Solid black lines represent results from loess regression, with gray areas representing
 4 95% confidence intervals.

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Fig. 2: Comparison of all of the ribosomal DNA copy number data from *Peronospora effusa* (dark gray, excluding zero values) obtained from the spore traps used in this study, and a lognormal distribution of the same mean and standard deviation (black outline), in 2013 and 2014.

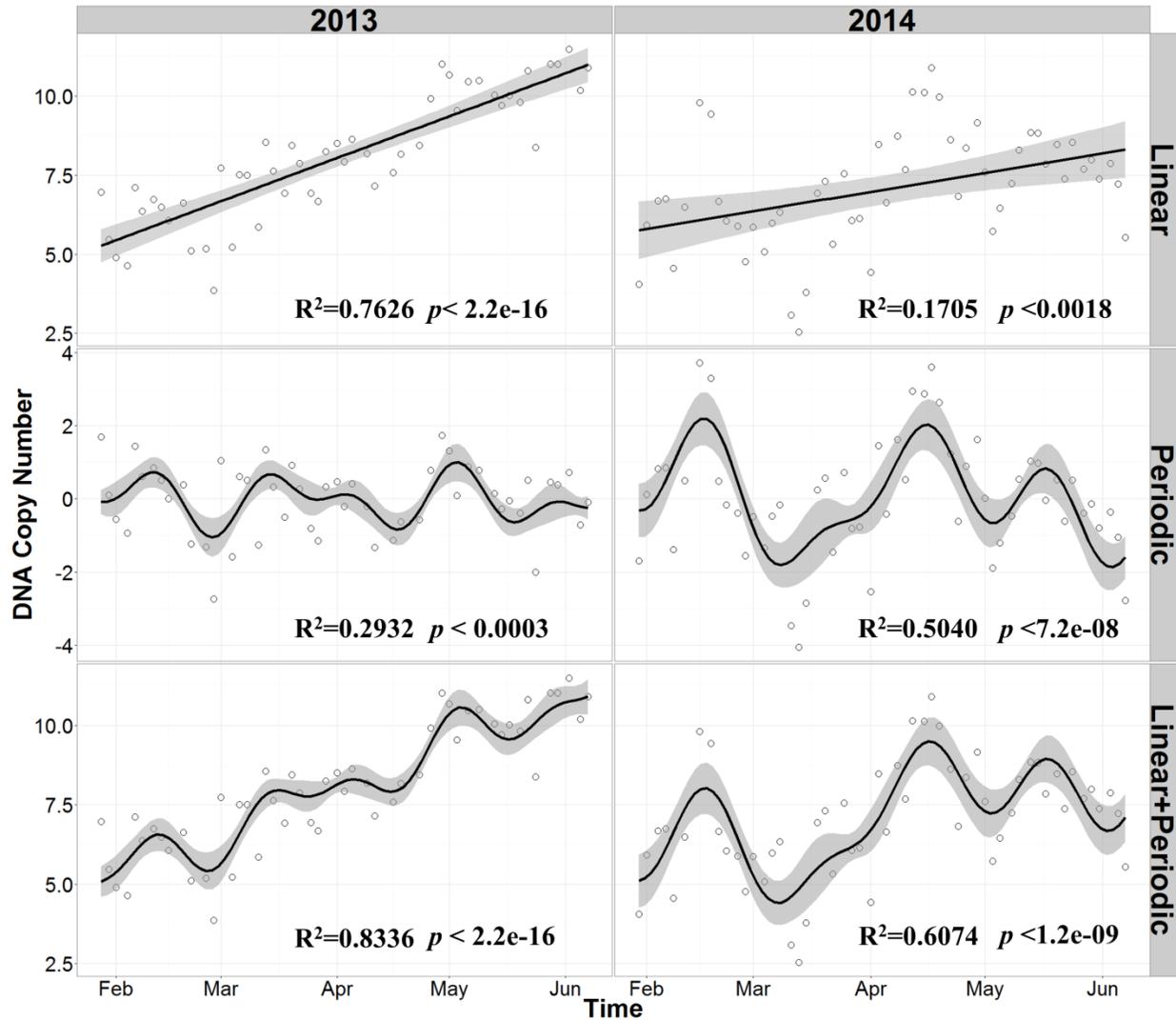


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2 **Fig. 3:** Scatterplot of the Lyapunov exponent and the coefficient of prediction from each of the
3 nine site-years.

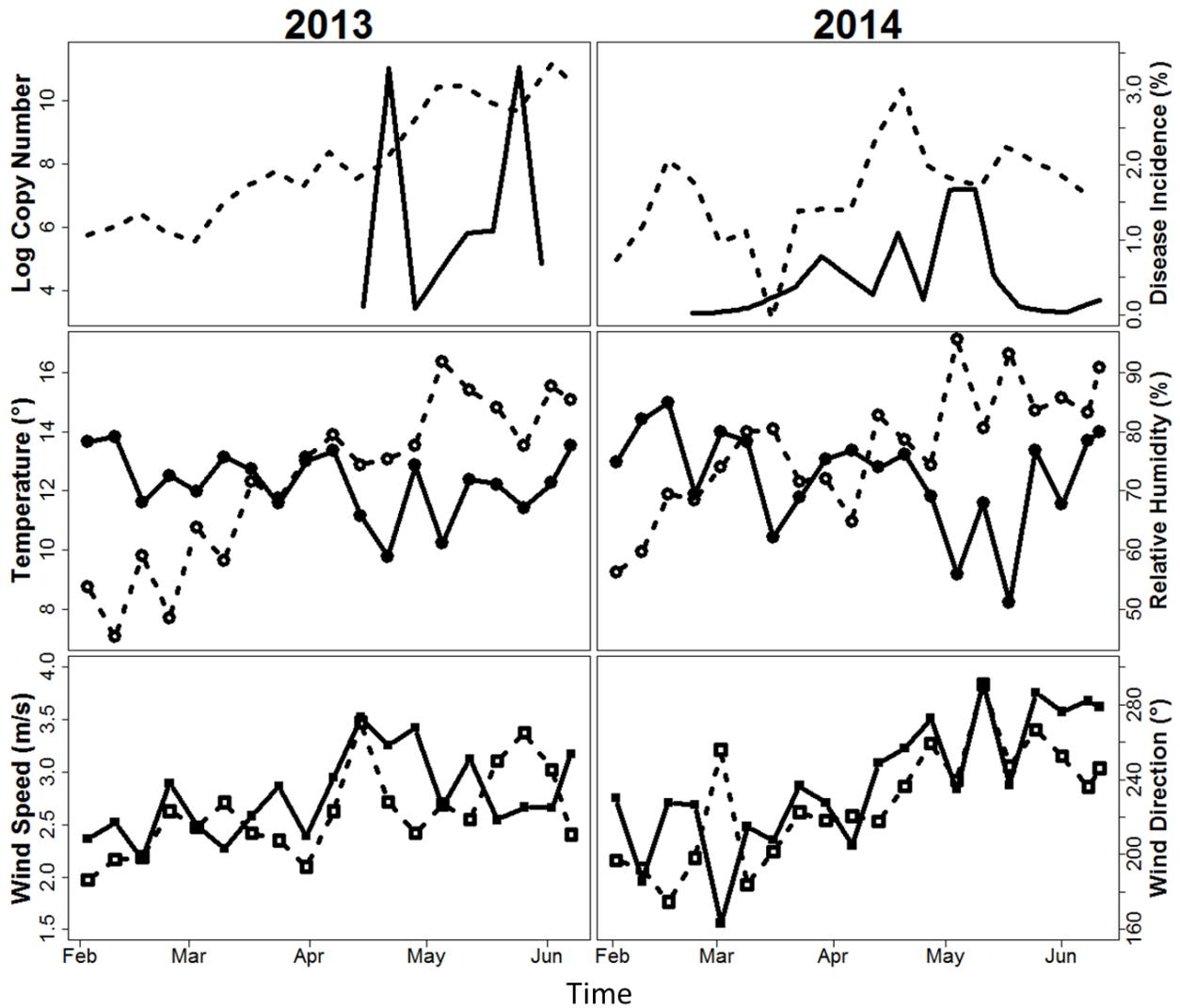
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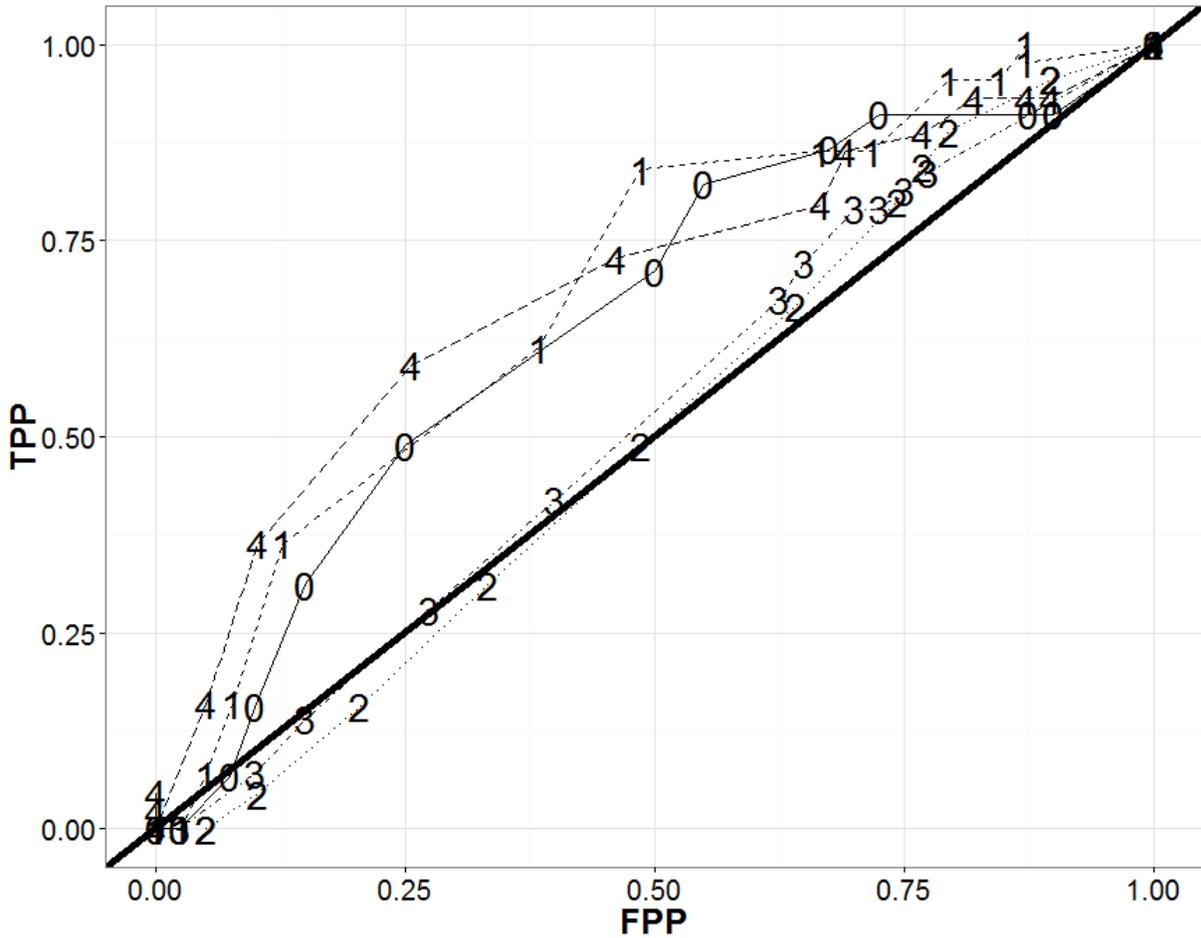


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 2 **Fig. 4:** Natural log of the ribosomal DNA copy number values from *Peronospora effusa*
 3 averaged daily across all sites in 2013 and 2014. Solid black lines represent results from
 4 regression analyses, with gray area representing 95% confidence intervals. The periodic and
 5 linear and periodic regressions were based on each year's period with the maximum coefficient
 6 of determination, which were 26 and 44 for 2013 and 30, 45, and 75 for 2014, respectively.

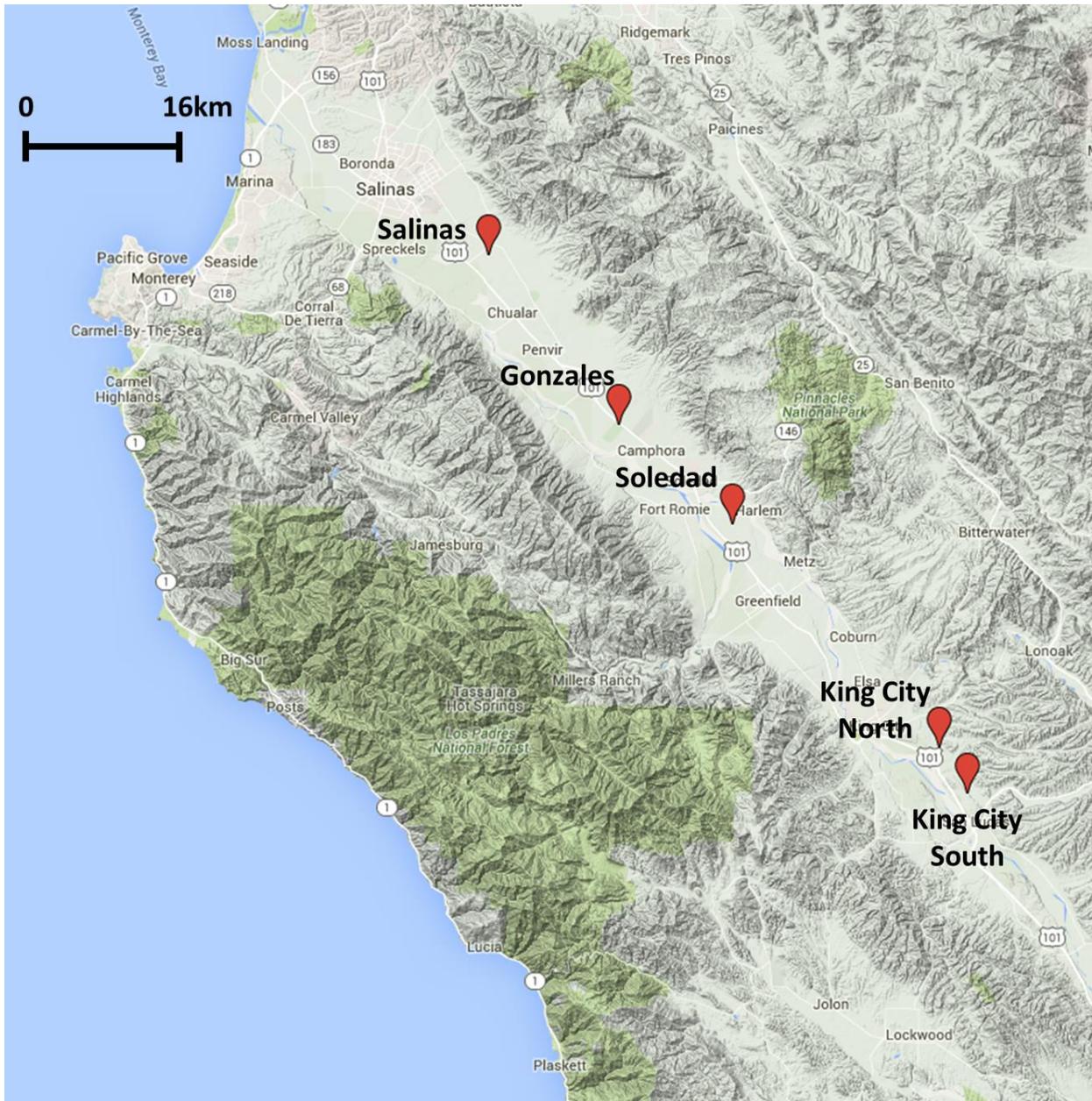
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 2 **Fig. 5:** Average weekly downy mildew disease and ribosomal DNA copy numbers from
 3 *Peronospora effusa* in 2013 and 2014. Dotted lines represent ribosomal DNA copy number and
 4 solid lines represent disease incidence ratings. Dotted lines with empty circles represent average
 5 weekly temperature and solid lines with solid circles represent relative humidity. Dotted lines
 6 with empty squares represent average weekly wind speed and solid lines with solid squares
 7 represent wind direction.

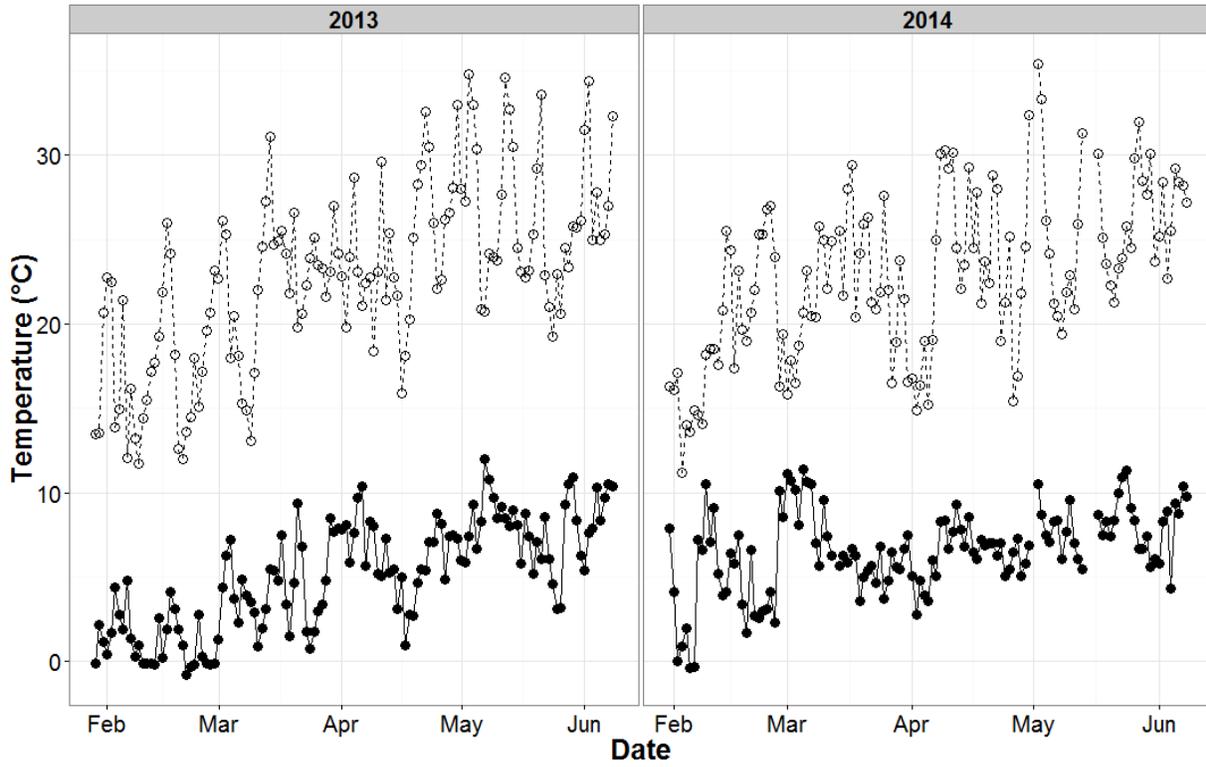


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 2 **Fig. 6:** Receiver operating characteristic (ROC) curve for different lags. Numbers indicate the
 3 lag period of the ribosomal DNA copy number from *Peronospora effusa* before the downy
 4 mildew disease incidence rating. Solid diagonal line represents the line of no discrimination.
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2 **Fig. S1:** Map of the spore trapping locations used in the study.

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 2 **Fig. S2:** Daily maximum and minimum temperatures across all sites in 2013 and 2014. Solid
 3 lines with solid circles represent minimum temperatures and dashed lines with hollow circles
 4 represent maximum temperatures.

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